

# Lab 10 – Repeated Measures

October 22 & 23, 2018  
FANR 6750

Richard Chandler and Bob Cooper

## Design

- We randomly assign each “subject” to a treatment
- We record the response to the treatment over time

## Sources of variation

- Treatment
- Time
- Treatment-time interaction
- Random variation among subjects
- Random variation within subjects

## APPROACHES

### Univariate

- This is just a split-plot analysis with adjusted  $p$ -values
- Adjustments: Greenhouse-Geisser or Huynh-Feldt methods
- In **R**, you must do a MANOVA to obtain these adjusted  $P$ -values

### MANOVA

- Testing based on Wilks' lambda or Pillai's trace
- This is usually followed by a profile analysis

### Mixed effects model with (ARMA) correlation structure

- This can be done using `lme`
- We might cover this later

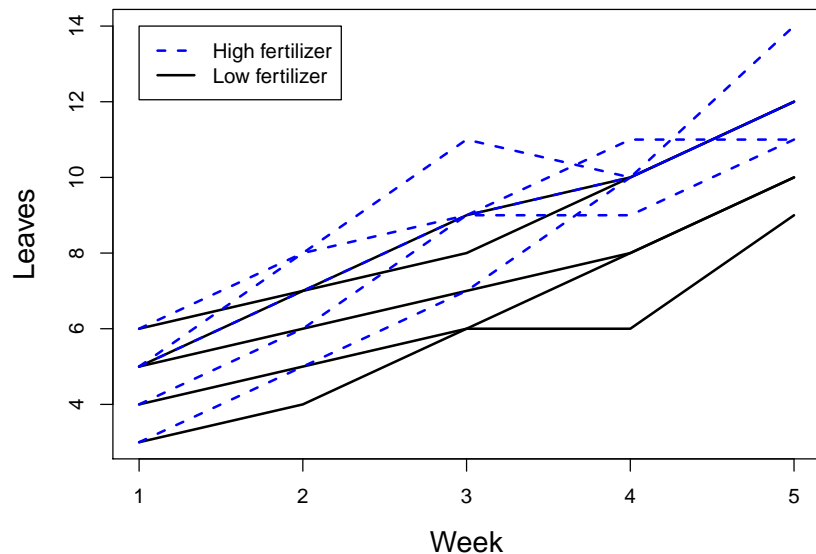
## THE PLANT DATA

```
plantData <- read.csv("plantData.csv")
plantData$plant <- factor(plantData$plant)
plantData$week <- factor(plantData$week)
str(plantData)

## 'data.frame': 50 obs. of 4 variables:
## $ plant      : Factor w/ 10 levels "1","2","3","4",...: 1 1 1 1 1 2 2 2 2 2 ...
## $ fertilizer: Factor w/ 2 levels "H","L": 2 2 2 2 2 2 2 2 2 2 ...
## $ week      : Factor w/ 5 levels "1","2","3","4",...: 1 2 3 4 5 1 2 3 4 5 ...
## $ leaves   : int  4 5 6 8 10 3 4 6 6 9 ...
```

```
head(plantData, n=8)

##   plant fertilizer week leaves
## 1     1           L     1      4
## 2     1           L     2      5
## 3     1           L     3      6
## 4     1           L     4      8
## 5     1           L     5     10
## 6     2           L     1      3
## 7     2           L     2      4
## 8     2           L     3      6
```



```
aov1 <- aov(leaves ~ fertilizer*week + Error(plant),
            data=plantData)
```

```
summary(aov1)
```

```
##
## Error: plant
##          Df Sum Sq Mean Sq F value Pr(>F)
## fertilizer 1  16.82   16.82   2.604  0.145
## Residuals  8   51.68    6.46
##
## Error: Within
##          Df Sum Sq Mean Sq F value Pr(>F)
## week      4  267.40   66.85 158.225 <2e-16 ***
## fertilizer:week 4   5.08    1.27   3.006  0.0326 *
## Residuals 32   13.52    0.42
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

We need to adjust the  $p$ -values for the time and interaction effects  
In R, this requires reformatting the data and running a MANOVA

```
plantData2 <- reshape(plantData, idvar="plant",
                    timevar="week", v.names="leaves",
                    direction="wide")
```

```
plantData2
##   plant fertilizer leaves.1 leaves.2 leaves.3 leaves.4 leaves.5
## 1     1           L         4         5         6         8         10
## 6     2           L         3         4         6         6         9
## 11    3           L         6         7         9         10        12
## 16    4           L         5         7         8         10        12
## 21    5           L         5         6         7         8         10
## 26    6           H         4         6         9         9         11
## 31    7           H         3         5         7         10        12
## 36    8           H         6         8         11        10        14
## 41    9           H         5         7         9         10        12
## 46   10           H         5         8         9         11        11
```

```
manova1 <- manova(cbind(leaves.1, leaves.2, leaves.3,
                       leaves.4, leaves.5) ~ fertilizer,
                 data=plantData2)
```

```
anova(manova1, X=~1, test="Spherical")
```

```
## Analysis of Variance Table
##
##
## Contrasts orthogonal to
## ~1
##
## Greenhouse-Geisser epsilon: 0.5882
## Huynh-Feldt epsilon: 0.8490
##
##          Df          F num Df den Df  Pr(>F)  G-G Pr  H-F Pr
## (Intercept) 1 158.2249    4    32 0.000000 0.000000 0.00000
## fertilizer   1   3.0059    4    32 0.032613 0.066622 0.04224
## Residuals    8
```

The last 3 columns are  $p$ -values corresponding to the effects of time (Intercept) and interaction (fertilizer). No adjustment is necessary for the main effect so you can use the  $p$ -value from `aov`.

Technically, adjusted  $p$ -values and MANOVA aren't necessary if the assumption of sphericity holds. **However**, we recommend doing the adjustments (or the MANOVA) anyway because the test of sphericity has low power.

Sphericity is the multivariate analogue of the homogeneity of variance assumption of ANOVA.

Here is how you test the assumption:

```
mauchly.test(manova1, X=~1)

##
## Mauchly's test of sphericity
## Contrasts orthogonal to
## ~1
##
##
## data:  SSD matrix from manova(cbind(leaves.1, leaves.2, leaves.3, leaves.4, lea
## W = 0.099297, p-value = 0.1062
```

We fail to reject the null hypothesis, so sphericity can be assumed.

```
summary(aov1)

##
## Error: plant
##          Df Sum Sq Mean Sq F value Pr(>F)
## fertilizer 1  16.82   16.82   2.604  0.145
## Residuals  8   51.68    6.46
##
## Error: Within
##          Df Sum Sq Mean Sq F value Pr(>F)
## week      4  267.40   66.85 158.225 <2e-16 ***
## fertilizer:week 4   5.08    1.27   3.006  0.0326 *
## Residuals 32  13.52    0.42
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
anova(manova1, X=~1, test="Spherical")

## Analysis of Variance Table
##
## Contrasts orthogonal to
## ~1
##
## Greenhouse-Geisser epsilon: 0.5882
## Huynh-Feldt epsilon:      0.8490
##
##          Df          F num Df den Df Pr(>F)  G-G Pr  H-F Pr
## (Intercept) 1 158.2249      4      32 0.000000 0.000000 0.000000
## fertilizer   1   3.0059      4      32 0.032613 0.066622 0.042224
## Residuals    8
```

## MULTIVARIATE TESTS – WILKS' LAMBDA

An alternative to the adjusted  $p$ -value approach is to do a multivariate analysis relaxing the assumptions about the structure of the variance-covariance matrix. We're already most of the way there.

```
anova(manova1, X=~1, test="Wilks")

## Analysis of Variance Table
##
##
## Contrasts orthogonal to
## ~1
##
##          Df      Wilks approx F num Df den Df      Pr(>F)
## (Intercept) 1 0.008487  146.042      4      5 2.308e-05 ***
## fertilizer   1 0.144772    7.384      4      5  0.02503 *
## Residuals    8
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

As before, we conclude that the effect of fertilizer changes over time

## MULTIVARIATE TESTS – PILLAI'S TRACE

Pillai's trace is an alternative to Wilks' lambda. In this case, it returns the same  $p$ -values as Wilks' test.

```
anova(manova1, X=~1, test="Pillai")

## Analysis of Variance Table
##
##
## Contrasts orthogonal to
## ~1
##
##          Df      Pillai approx F num Df den Df      Pr(>F)
## (Intercept) 1 0.99151  146.042      4      5 2.308e-05 ***
## fertilizer   1 0.85523    7.384      4      5  0.02503 *
## Residuals    8
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Profile analysis requires calculating the differences (ie, the number of leaves grown each week).

```
manova2 <- manova(
  cbind(leaves.2-leaves.1, leaves.3-leaves.2,
        leaves.4-leaves.3, leaves.5-leaves.4) ~
  fertilizer, data=plantData2)
```

During which intervals do the growth rates differ?

```
summary.aov(manova2)

## Response 1 :
##           Df Sum Sq Mean Sq F value Pr(>F)
## fertilizer  1   2.5     2.5    12.5 0.00767 **
## Residuals   8   1.6     0.2
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Response 2 :
##           Df Sum Sq Mean Sq F value Pr(>F)
## fertilizer  1   1.6     1.6     3.2 0.1114
## Residuals   8   4.0     0.5
##
## Response 3 :
##           Df Sum Sq Mean Sq F value Pr(>F)
## fertilizer  1   0.1     0.1    0.0625 0.8089
## Residuals   8  12.8     1.6
##
## Response 4 :
##           Df Sum Sq Mean Sq F value Pr(>F)
## fertilizer  1   0.1     0.1    0.0909 0.7707
## Residuals   8   8.8     1.1
```

## PLOT THE GROWTH RATES

Calculate mean growth rate for each time interval

```
leavesMat <- plantData2[,3:7]
growthMat <- leavesMat[,2:5] - leavesMat[,1:4]
colnames(growthMat) <- paste("interval", 1:4, sep=".")
(lowFertilizer <- colMeans(growthMat[1:5,]))

## interval.1 interval.2 interval.3 interval.4
##          1.2          1.4          1.2          2.2

(highFertilizer <- colMeans(growthMat[6:10,]))

## interval.1 interval.2 interval.3 interval.4
##          2.2          2.2          1.0          2.0
```

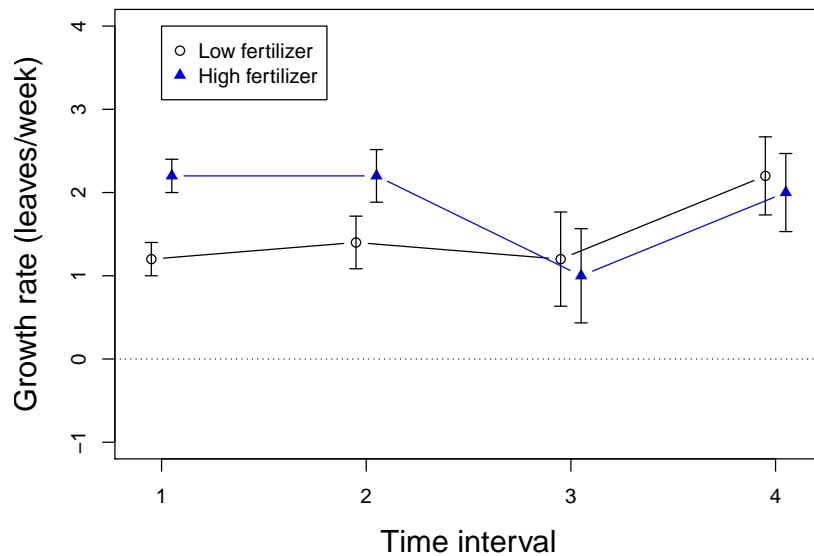
Calculate the standard errors for these growth rates

```
SE <- sqrt(diag(stats::vcov.mlm(manova2)))
SE <- SE[names(SE)==":(Intercept)"] # Only use "intercept" SEs
unnname(SE) ## Ignore the names

## [1] 0.2000000 0.3162278 0.5656854 0.4690416
```

## PLOT THE GROWTH RATES

```
plot(1:4-0.05, lowFertilizer, type="b", xlim=c(0.9, 4.1),
     ylim=c(-1, 4), xaxp=c(1,4,3), cex.lab=1.5,
     xlab="Time interval", ylab="Growth rate (leaves/week)")
abline(h=0, lty=3)
arrows(1:4-0.05, lowFertilizer-SE, 1:4-0.05, lowFertilizer+SE,
       angle=90, code=3, length=0.05)
lines(1:4+0.05, highFertilizer, type="b", pch=17, col=4)
arrows(1:4+0.05, highFertilizer-SE, 1:4+0.05,
       highFertilizer+SE, angle=90, code=3, length=0.05)
legend(1, 4, c("Low fertilizer", "High fertilizer"),
      col=c("black", "blue"), pch=c(1,17))
```



A researcher wants to assess the effects of crowding on the growth of the dark toadfish (*Neophrynichthys latus*). 15 fish tanks are stocked with three densities of conspecifics. Five tanks have low density (1 fish), 5 tanks have medium density (5 fish), and 5 tanks have high density (10 fish). In each tank, the weight of one "focal fish" is recorded on 6 consecutive weeks. The data are in the file `fishData.csv`.

- (1) Conduct the univariate repeated measures ANOVA using `aov`. Calculate the adjusted  $p$ -values using the Huynh-Feldt method. Does the effect of density change over time?
- (2) Conduct a multivariate repeated measures ANOVA and use Wilks' lambda to test if the effect of density changes over time. What is your conclusion?
- (3) Conduct a profile analysis. In which time intervals is the effect of density on growth rate significant?

Upload your self-contained R-script to ELC at least one day before your next lab